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# Diagnostic value of cerebrospinal fluid Aß ratios in preclinical Alzheimer's disease

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#### Abstract

**Introduction:** In this study of preclinical Alzheimer's disease (AD) we assessed the added diagnostic value of using cerebrospinal fluid (CSF) Aβ ratios rather than Aβ42 in isolation for detecting individuals who are positive on amyloid positron emission tomography (PET).

**Methods:** Thirty-eight community-recruited cognitively intact older adults (mean age 73, range 65–80 years) underwent <sup>18</sup>F-flutemetamol PET and CSF measurement of A $\beta$ 1-42, A $\beta$ 1-40, A $\beta$ 1-38, and total tau (ttau). <sup>18</sup>F-flutemetamol retention was quantified using standardized uptake value ratios in a composite cortical region (SUVR<sub>comp</sub>) with reference to cerebellar grey matter. Based on a prior autopsy validation study, the SUVR<sub>comp</sub> cut-off was 1.57. Sensitivities, specificities and cut-offs were defined based on receiver operating characteristic analysis with CSF analytes as variables of interest and <sup>18</sup>F-flutemetamol positivity as the classifier. We also determined sensitivities and CSF cut-off values at fixed specificities of 90 % and 95 %.

**Results:** Seven out of 38 subjects (18 %) were positive on amyloid PET. A $\beta$ 42/tau, A $\beta$ 42/A $\beta$ 40, A $\beta$ 42/A $\beta$ 38, and A $\beta$ 42 had the highest accuracy to identify amyloid-positive subjects (area under the curve (AUC)  $\geq$  0.908). A $\beta$ 40 and A $\beta$ 38 had significantly lower discriminative power (AUC = 0.571). When specificity was fixed at 90 % and 95 %, A $\beta$ 42/ttau had the highest sensitivity among the different CSF markers (85.71 % and 71.43 %, respectively). Sensitivity of A $\beta$ 42 alone was significantly lower under these conditions (57.14 % and 42.86 %, respectively).

**Conclusion:** For the CSF-based definition of preclinical AD, if a high specificity is required, our data support the use of Aβ42/ttau rather than using Aβ42 in isolation.

#### Introduction

Preclinical [1, 2], or asymptomatic [3], Alzheimer's disease (AD) is characterized by the presence of AD-related pathophysiological processes in the absence of cognitive deficits. Evidence of brain amyloidosis is a requirement common to all three National Institute on Ageing–Alzheimer's Association (NIA–AA) stages of preclinical AD [1] and is also a defining feature of the asymptomatic at risk for AD state according to the International Working Group IWG-2 criteria [3]. This can be detected directly in vivo by means of either amyloid-beta (A $\beta$ ) protein quantification in



Apart from A $\beta$ 1–42, other A $\beta$  isoforms (e.g., A $\beta$ 1–40, A $\beta$ 1–38) have evoked interest from a clinical-diagnostic perspective, as either a separate biomarker tool or when combined (ratio) with A $\beta$ 1–42 [6–8]. Using ratios of A $\beta$  isoforms (A $\beta$ 1–42/A $\beta$ 1–38, A $\beta$ 1–42/A $\beta$ 1–40) may have added value for the discrimination between AD and normal pressure hydrocephalus [9], cerebral amyloid angiopathy [10], frontotemporal dementia [11], and Lewy body dementia [12], and also between mild cognitive impairment (MCI) due to AD versus non-AD MCI [13]. In cognitively intact individuals, A $\beta$ 38 or A $\beta$ 40 does not correlate with amyloid PET positivity, in contrast with A $\beta$ 42 [5, 14].

In this study of preclinical AD, we assessed the added value of using ratios of  $A\beta 42$  to other C-terminal  $A\beta$  isoforms or to total tau (ttau) for discriminating amyloid-



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positive versus amyloid-negative cognitively intact healthy controls, with an autopsy-validated <sup>18</sup>F-flutemetamol cutoff score [15] as standard of truth. The cutoff value was derived from the <sup>18</sup>F-flutemetamol phase 3 study using a binarized measure of postmortem brain neuritic plaque density [16] (overall mean Bielschowsky score below or above 1.5 [15]). We also explored the diagnostic value of the Aβ38 and Aβ40 isoforms on their own.

For design of clinical trials in preclinical AD, the data presented may inform the decision on which CSF parameter to select for study eligibility based on its equivalence to an amyloid-PET-based definition. We not only provide the parameters providing optimal balance between sensitivity and specificity but also the parameters that provide an acceptable sensitivity for a fixed high specificity. Specificity may receive more weight in trials in preclinical AD because the definition of the target population often heavily relies on the biomarker value, healthy volunteers are exposed to potential adverse effects of study drugs for a long duration, and positive evidence for the presence of the study target increases the likelihood of success. Sensitivity will mainly determine the number needed to screen, and will therefore impact on the cost.

#### Methods

#### Participants

Thirty-eight cognitively intact older controls (mean age 73 years, standard deviation (SD) 5 years; Table 1) were recruited prospectively and consecutively, from 10 September 2012 until 4 April 2014, through advertisement in local newspapers and through websites for seniors, asking for healthy volunteers between 65 and 80 years of age for participation in a scientific study at the University Hospital Leuven, Belgium, involving brain imaging (sic). At screening, subjects underwent a detailed interview

**Table 1** Demographics and CSF biomarker concentrations

about medical history, a Mini Mental State Examination (MMSE), a Clinical Dementia Rating (CDR), blood sampling, and a conventional neuropsychological assessment. Inclusion criteria were age 65–80 years, MMSE  $\geq 27$ , CDR = 0, and normal test scores on neuropsychological assessment according to the published norms adapted for age, gender, and education. Among the exclusion criteria were a neurological or psychiatric history and focal brain lesions on structural magnetic resonance imaging (MRI). Subjects who fulfilled all criteria underwent both <sup>18</sup>F-flute-metamol PET and lumbar puncture. The target sample size of the PET-plus-CSF cohort was 40 but two subjects dropped out after the PET scan and prior to the lumbar puncture, giving a final sample size of 38.

This PET-plus-CSF cohort belonged to a larger cohort of healthy older controls undergoing <sup>18</sup>F-flutemetamol PET (target sample n = 180, recruited until time of writing n = 172) [17, 18]. The other subjects in this larger cohort did not undergo lumbar puncture per protocol. The primary aim of the full cohort was to investigate the interaction between brain-derived neurotrophic factor (BDNF) and apolipoprotein E (APOE) genetic polymorphisms on amyloid deposition and functional reorganization [17, 18]. The inclusion and exclusion criteria for the full cohort were identical to those of the PET-plus-CSF cohort apart from the age range (50-80 years for the full cohort). At inclusion, participants of the full cohort were stratified per age bin for two genetic factors: BDNF (met allele at codon 66 present or absent) and APOE (E4 allele present or absent). The cells of this  $2 \times 2$  factorial design were prospectively matched for number of cases, APOE and BDNF genetic status, age, sex, and education.

The PET-plus-CSF cohort (n = 38) did not differ from the remaining subjects (n = 134) with regards to sex, education, number of APOE  $\varepsilon 4$  carriers or BDNF *met* carriers, the presence of subjective memory complaints

Table T Demographics and CSF biomarker concentrations							
Gender (male/female)	22/16	LVF (number of words)	36.0 (10.8, 17–64)				
Age (years)	73 (4.7, 65–80)	RPM (/60)	36.1 (9.8, 15–53)				
Education (years)	13.4 (3.1, 8–20)	TMT B/A	2.4 (0.5, 1.5–3.8)				
APOE ɛ4 carriers (n)	19 (50 %)	Aβ38 (pg/ml)	2401 (654, 1057–3505)				
BDNF met carriers (n)	20 (53 %)	Aβ40 (pg/ml)	8933 (2456, 3640–13273)				
MMSE (/30)	28.9 (1.0, 27–30)	Aβ42 (pg/ml)	996 (430, 351–1859)				
AVLT TL (/75)	46.2 (8.4, 31–69)	ttau (pg/ml)	360 (134, 126–660)				
AVLT DR (/15)	9.8 (2.5, 5–14)	Αβ42/Αβ38	0.412 (0.119, 0.136–0.596)				
AVLT %DR	83.7 (11.7, 55–108)	Αβ42/Αβ40	0.110 (0.030, 0.044–0.148)				
BNT (/60)	54.2 (4.2, 41–60)	Aβ42/ttau	3.015 (1.246, 0.749–5.128)				
AVF (number of words)	24.0 (5.5, 14–40)	Amyloid+ (n)	7 (18 %)				

Data presented as mean (standard deviation, range)

Aβ amyloid beta, APOE apolipoprotein E, AVF Animal Verbal Fluency Test, AVLT Rey Auditory Verbal Learning Test, BDNF brain-derived neurotrophic factor, BNT Boston Naming Test, CSF cerebrospinal fluid, DR delayed recall, LVF Letter Verbal Fluency Test, MMSE Mini Mental State Examination, RPM Raven's Progressive Matrices, TL total learning, TMT Trail Making Test (part B divided by part A), ttau total tau (29 % in each of the two groups), or neuropsychological test scores (P > 0.23). The CSF cohort was significantly older than the remaining subjects (mean age 73 years vs. mean age 67 years, P < 0.0001). The proportion of amyloid-positive cases did not differ significantly between the CSF-plus-PET cohort (18 %) and the remaining subjects (12 %) (P = 0.23).

The protocol (EudraCT: 2009-014475-45) was approved by the Ethics Committee University Hospitals Leuven, Belgium. Written informed consent was obtained from all subjects in accordance with the Declaration of Helsinki.

#### Amyloid PET

<sup>18</sup>F-flutemetamol PET was acquired on a 16-slice Siemens Biograph PET/CT scanner (Siemens, Erlangen, Germany). The tracer was injected as a bolus into an antecubital vein (mean activity 150 MBq, SD 5 MBq, range 134-162 MBq). Scan acquisition started 90 minutes after tracer injection and lasted for 30 minutes [17-20]. Prior to PET acquisition, a low-dose computed tomography scan of the head was performed for attenuation correction. Random and scatter correction were applied. The PET summed image was spatially normalized to Montreal Neurological Institute (MNI) space using a fully automated PET-only method [21]. On the basis of spatially normalized images (voxel size  $2 \times 2 \times 2$  mm<sup>3</sup>), standardized uptake value ratios (SUVR) were calculated with cerebellar gray matter as the reference region. The mean SUVR value was calculated in a composite cortical region  $(SUVR_{comp})$  [15]. The composite cortical region and the cerebellar gray matter reference region were defined as a combination of narrow automated anatomic labeling-type regions [22] outlined on the ICBM-152 template masked with a gray matter probability mask [15]. Images were analyzed by an experienced medical imaging specialist blinded to all study information.

To estimate the SUVR<sub>comp</sub> cutoff value for detecting amyloid positivity in vivo using the described method, receiver operating curve (ROC) analysis was performed by Thurfjell et al. [15] on an independent dataset of 68 SUVR<sub>comp</sub> values (quantified based on the already described method) with the autopsy results as a standard of truth. The autopsy data were classified following Vemuri's modification of the Consortium to Establish a Registry for AD criteria [16, 23]. Eight cortical regions (precuneus, midfrontal cortex, superior temporal cortex, middle temporal cortex, inferior parietal cortex, anterior cingulate gyrus, posterior cingulate gyrus, and primary visual cortex) were scored using an overall mean Bielschowsky score: 0 = no plaques, 1 = one to five plaques, 2 = six to 19 plaques, 3 = 20 or more plaques. If the mean Bielschowsky score was > 1.5 in at least one region, the brain was classified as amyloid-positive; if all regions scored  $\leq$  1.5, the brain was classified as amyloidnegative. The resulting  ${\rm SUVR}_{\rm comp}$  cutoff value was 1.57 [15].

#### Lumbar puncture and CSF analysis

Lumbar punctures were carried out at the L4/5 level in the morning (10 a.m.-2 p.m.) and collected in polypropylene tubes (total volume 15 ml, Greiner Bio-one Cellstar; VWR, Leuven, Belgium), discarding 1 ml to avoid traumatic blood contamination. Samples were centrifuged within 30 minutes after collection (2600 rpm, 10 minutes, 4 °C). After centrifugation, supernatants were transferred into polypropylene tubes and from there aliquoted in 1.5 ml polypropylene tubes (1 ml volume CSF/tube; Kartell, Noviglio, Italy). Samples were stored at -80 °C until batch analysis. Our primary analysis was based on the EUROIMMUN single analyte enzyme-linked immunosorbent assays (ELISA) (EUROIMMUN, Lübeck, Germany) of CSF A<sub>β1-42</sub>, A<sub>β1-40</sub>, A<sub>β1-38</sub>, and ttau. The assays were performed at ADx Ghent, Belgium by two experienced laboratory technicians blinded to all study information. The A $\beta$  assays quantify the full length of the C-terminus-specific  $A\beta$  isoforms ( $A\beta$ 1-specific assay format). The tau assay is designed with a capture antibody towards the central region and one monoclonal antibody with an epitope at the amino-terminus of the protein. The assay design includes lyophylized recombinant proteins as calibrators, run-validation control samples (calibrators added to a phosphate-buffered solution), as well as a qualification panel to evaluate the analytical performance(s) in the laboratory. These novel immunoassays are free from matrix interference and their intra-assay reproducibility has a coefficient of variation  $\leq 5.0$  % with an inter-assay reproducibility  $\leq 8.3 \%$  [24].

As a secondary analysis, we verified our results using the INNOTEST ELISA for  $A\beta 1-42$ , ttau, and <sup>181</sup>phospho-tau (ptau) (Fujirebio Europe, Ghent, Belgium). The assays were performed at the Laboratory Medicine Department of UZ Leuven, Belgium, in a ISO-15189 and Joint Commission International accredited environment by an expert technician blinded to all study information. The assay design included ready-to-use recombinant proteins as calibrators, run-validation control samples, and internal quality controls samples (for which target value and acceptance criteria were established in the routine setting of AD biomarker quantification).

#### Statistical analysis

In the primary analysis, which was based on the EURO-IMMUN assays, we compared the diagnostic accuracy of different CSF A $\beta$  isoforms, their ratios, ttau, and A $\beta$ 42/ ttau to detect amyloid-positive older individuals. We used a ROC analysis with CSF analytes as variables of interest and <sup>18</sup>F-flutemetamol positivity defined based on the autopsy-derived SUVR<sub>comp</sub> cutoff value as a classifier.

We also evaluated whether case classification changed when we varied the cutoff value by  $\pm 1.5$  %, corresponding to the test–retest variability estimated for SUVR<sub>comp</sub> [20]. The highest Youden index (sensitivity + specificity – 1) was used to estimate the optimal ROC cutoff values. Statistical differences between ROCs were evaluated according to the method of DeLong et al. [25] for pairwise ROC comparisons. Correction for multiple comparisons (n = 21) was performed with the Bonferroni method. The Bonferroni corrected threshold for significance was P < 0.002, corresponding to  $P_{corrected} < 0.05$ .

Depending on the study, a high specificity may be desirable even if this implies a loss of sensitivity. We therefore also evaluated sensitivities and cutoff values at a fixed prespecified specificity of 90 % and 95 %, respectively. We evaluated whether this changed case classification significantly (McNemar test).

As a secondary analysis, we performed ROC analyses based on the INNOTEST assay of  $A\beta_{42}$ , ttau, and ptau and statistically compared the areas under the curves (AUCs) between the two types of assays. We also compared the AUCs between the different INNOT-EST measures and determined the sensitivity and percentage of correct classifications at a fixed specificity of 90 % and 95 %.

As a further secondary analysis, we evaluated the continuous relationship between the different CSF analytes and  $^{18}\rm{F}$ -flutemetamol SUVR<sub>comp</sub> values. We tested

whether a linear, polynomial (quadratic), exponential, or hyperbolic relation fitted best to these data. The model assumptions were assessed by evaluating normality and homoscedasticity of residuals with q–q plots and plots of residuals versus fitted values. The best fitting model was selected based on the Akaike information criterion (AIC), which is a measure of model fit. A lower AIC indicates a better fit. CSF analytes were used as dependent variables and <sup>18</sup>F-flutemetamol SUVR<sub>comp</sub> as an independent variable.

Statistical analyses were performed in R version 3.1.1 (https://www.r-project.org) and MedCalc version 14.8.1 (https://www.medcalc.org).

#### Results

Based on the autopsy-confirmed  $^{18}\rm F$ -flutemetamol SUVR<sub>comp</sub> cutoff value, seven out of 38 subjects (18 %) were assigned to the amyloid-positive category (Fig. 1a). Case assignment did not change when we varied the cut-off value according to the known test–retest replicability.

APOE  $\varepsilon$ 4 carriers had significantly lower values of A $\beta$ 42, A $\beta$ 42/ttau, A $\beta$ 42/A $\beta$ 40, and A $\beta$ 42/A $\beta$ 38 than  $\varepsilon$ 4 noncarriers (*P* < 0.003). CSF analyte concentrations did not differ between BDNF *met* carriers and noncarriers (*P* > 0.23).

Aβ42/ttau, Aβ42/Aβ40, Aβ42/Aβ38, and Aβ42 discriminated between <sup>18</sup>F-flutemetamol-positive and <sup>18</sup>F-flutemetamol-negative subjects with high accuracy (AUC  $\geq$  0.908;





Table 2, Fig. 1b). A $\beta$ 38, A $\beta$ 40, and ttau showed a lower discriminative power with AUC  $\leq$  0.724 (Table 2). A $\beta$ 42/ ttau, A $\beta$ 42/A $\beta$ 40, and A $\beta$ 42 had significantly higher AUCs than A $\beta$ 38 or A $\beta$ 40 alone (Table 2, *P* < 0.003). A $\beta$ 42/A $\beta$ 38 had significantly higher AUCs than A $\beta$ 40 (*P* = 0.002). There was no significant difference between the ratios A $\beta$ 42/ttau, A $\beta$ 42/A $\beta$ 40, and A $\beta$ 42/A $\beta$ 38, on the one hand, and A $\beta$ 42 alone, on the other (Table 2, *P* > 0.32). The AUCs of the three ratios were not statistically different from each other (Table 2, *P* > 0.30).

When specificity was fixed at 90 %,  $A\beta 42/tau$  and  $A\beta 42/A\beta 40$  had the highest sensitivity and  $A\beta 42/A\beta 38$  the second highest sensitivity (Table 3). All three  $A\beta$  isoforms ( $A\beta 42$ ,  $A\beta 40$ ,  $A\beta 38$ ) used on their own detected significantly fewer amyloid PET-positive cases when specificity was fixed a priori at 90 % than when the cutoff value was based on the highest Youden index (Table 3), indicative of a significant loss in sensitivity. This was not the case for  $A\beta 42/tau$ ,  $A\beta 42/A\beta 40$ , and  $A\beta 42/A\beta 38$  ratios and ttau (Table 3).

When specificity was fixed at 95 %, A $\beta$ 42/ttau had the highest sensitivity (Table 3). All A $\beta$  isoforms, ttau, and all ratios detected significantly less amyloid-positive cases when the specificity was fixed a priori at 95 % compared with the highest Youden index-based cutoff value, with one exception—namely the ratio A $\beta$ 42/ttau (Table 3). At a specificity of 95 %, the number of amyloid PET-positive cases detected based on the ratio A $\beta$ 42/ttau did not differ significantly from the number detected based on the highest Youden index-based cutoff value, although it was numerically lower.

As a secondary analysis, we compared the AUCs between two types of assays, EUROIMMUN and INNOT-EST. The AUCs for A $\beta$ 42, ttau, and A $\beta$ 42/ttau did not differ between the EUROIMMUN and INNOTEST assays (A $\beta$ 42, P = 0.33; ttau, P = 0.91; and A $\beta$ 42/ttau, P = 0.25) (Tables 2 vs. 4). When we compared the AUCs between the different INNOTEST measures, the AUC for A $\beta$ 42/ ttau differed significantly from the AUC for ttau (uncorrected *P* = 0.0172) or ptau (uncorrected *P* = 0.0096). When specificity was fixed at 90 %, A $\beta$ 42 and A $\beta$ 42/ttau had the highest sensitivity (Table 4). When specificity was fixed at 95 %, A $\beta$ 42/ttau had the highest sensitivity (Table 4).

Four CSF analytes—A $\beta$ 42/ttau, A $\beta$ 42/A $\beta$ 40, A $\beta$ 42/A $\beta$ 38, and A $\beta$ 42—showed a significant correlation with the <sup>18</sup>F-flutemetamol SUVR<sub>comp</sub> values (Fig. 2). The linear model was rejected because it did not satisfy assumptions of the model. The hyperbolic model fitted best to the relationship between A $\beta$ 42 and <sup>18</sup>F-flutemetamol SUVR<sub>comp</sub>. The relationships between <sup>18</sup>F-flutemetamol SUVR<sub>comp</sub> and A $\beta$ 42/ttau, A $\beta$ 42/A $\beta$ 40, and A $\beta$ 42/A $\beta$ 38 were best described by the exponential model. However, differences between the models were small. There was no correlation between <sup>18</sup>F-flutemetamol SUVR<sub>comp</sub> values and A $\beta$ 38, A $\beta$ 40, and ttau (Fig. 2).

#### Discussion

Overall, when sensitivity and specificity were combined, the ability to discriminate amyloid-positive from amyloidnegative cognitively healthy older adults was comparable between A $\beta$ 42 on its own and the ratio of A $\beta$ 42 over the isoforms examined or over ttau. However, when a high specificity of 90-95 % was imposed as a criterion, the sensitivity of A $\beta$ 42 alone diminished to 43–57 %. The sensitivity of the ratio over A $\beta$ 40 was acceptable at a specificity of 90 % (86 %), but at a specificity of 95 % the sensitivity decreased to 57 %. Under these requirements, the ratio over ttau was the only measure which retained an acceptable sensitivity (71-86 %). A high specificity would for instance be desirable if the potential benefit of a study drug depends on the amyloid positivity of cognitively normal subjects and the study drug has potentially noxious effects or a high cost. A favorable tradeoff in terms of sensitivity, as was the case only for A $\beta$ 42 over ttau, would decrease

**Table 2** Diagnostic performance of different CSF analytes with <sup>18</sup>F-flutemetamol PET as autopsy-validated standard of truth (EUROIMMUN assay)

	AUC	SE	95 % CI	Cutoff <sup>a</sup>	Sensitivity (%)	Specificity (%)	Correctly classified <sup>b</sup> (%)
Αβ38	0.571	0.111	0.401-0.730	2909	100	32.26	45
Αβ40	0.571	0.112	0.401-0.730	10738	100	29.03	42
Αβ42*†	0.908	0.051	0.769–0.977	745	100	74.19	79
ttau	0.724	0.148	0.555-0.856	436	71.43	80.65	76
Αβ42/Αβ38*	0.935	0.039	0.806-0.989	0.332	100	87.10	89
Αβ42/Αβ40*†	0.954	0.033	0.832-0.995	0.096	100	80.65	84
Aβ42/ttau*†	0.963	0.028	0.846-0.998	2.006	100	87.10	89

Analyte concentrations are described as pg/ml or calculated as ratios between concentrations of two analytes

Statistically significant differences of AUCs between analytes:  $*P_{corrected} < 0.05$  compared with A $\beta$ 40;  $†P_{corrected} < 0.05$  compared with A $\beta$ 38. No other differences of AUCs were found

<sup>a</sup>Cutoff value corresponding to the highest Youden index

<sup>b</sup>Percentage of positively classified cases based on the CSF cutoff compared with amyloid PET classification

Aβ amyloid beta, AUC area under the receiver operating characteristic curve, Cl confidence interval, CSF cerebrospinal fluid, PET positron emission tomography, SE standard error, ttau total tau

	Sensitivity (%)	95 % CI	Cutoff value	Difference <sup>a</sup> (%)	P value <sup>b</sup>	Correctly classified <sup>c</sup> (%)
Specificity of 90 %						
Αβ38	14.29	0.00-71.43	1446	65.79	<0.0001	79
Αβ40	14.29	0.00-71.43	5602	65.79	<0.0001	76
Αβ42	57.14	0.00-100.00	546	21.05	0.008	84
ttau	57.14	14.29–100.00	471	10.53	0.125	82
Αβ42/Αβ38	71.43	0.00-100.00	0.268	7.89	0.250	87
Αβ42/Αβ40	85.71	14.29–100.00	0.074	10.53	0.125	89
Aβ42/ttau	85.71	14.29–100.00	1.852	5.26	0.500	89
Specificity of 95 %						
Αβ38	14.29	0.00-71.43	1342	68.42	<0.0001	82
Αβ40	14.29	0.00-71.43	5254	71.05	<0.0001	82
Αβ42	42.86	0.00-85.71	493	28.95	0.001	87
ttau	42.86	0.00-85.71	539	18.42	0.016	84
Αβ42/Αβ38	28.57	0.00-71.43	0.251	21.05	0.008	84
Αβ42/Αβ40	57.14	8.62-85.71	0.067	21.05	0.008	89
Aβ42/ttau	71.43	28.57-100.00	1.415	13.16	0.063	92

Table 3 Clinical accuracy: estimated sensitivities and cutoff values at a fixed specificity of 90 % or 95 % (EUROIMMUN assay)

Analyte concentrations are described as pg/ml or calculated as ratios between concentrations of two analytes

<sup>a</sup>Percentage of subjects who were classified differently based on the cutoff values from fixed specificities compared with the cutoff values corresponding to the highest Youden index (Table 2)

<sup>b</sup>Significance for the "Difference"<sup>C</sup>Percentage of positively classified cases based on the CSF cutoffs from fixed specificities compared with amyloid PET classification

Aβ amyloid beta, Cl confidence interval, ttau total tau

Table 4 Diagnostic performance of different CSF	alytes measured with the INNOTEST assay for A $eta$ 42, ttau, and ptau at an optim	al
specificity and at a specificity fixed at 90 % or 95		

	AUC	SE	95 % CI	Cutoff <sup>a</sup>	Sensitivity (%)	Specificity (%)	Correctly classified <sup>b</sup> (%)
Αβ42	0.935	0.0394	0.806-0.989	853	100	83.87	87
ttau	0.733	0.132	0.565–0.863	352	71.43	77.42	76
ptau	0.675	0.139	0.504-0.818	86	42.86	93.55	84
Aβ42/ttau	0.880	0.0878	0.734-0.963	2.258	85.71	90.32	89
Specificity of 90 %	Sensitivit	ty (%)	95 % CI	Cutoff <sup>a</sup>	Difference <sup>c</sup> (%)	P value <sup>d</sup>	Correctly classified <sup>b</sup> (%)
Αβ42	85.71		11.54-100.00	798	7.90	0.25	89
ttau	57.14		14.29-100.00	465	10.53	0.125	82
ptau	42.96		0.00-85.71	87	5.26	0.5	79
Aβ42/ttau	85.71		28.57-100.00	2.263	0	1	89
Specificity of 95 %	Sensitivit	ty (%)	95 % CI	Cutoff <sup>a</sup>	Difference <sup>c</sup> (%)	P value <sup>d</sup>	Correctly classified <sup>b</sup> (%)
Αβ42	42.86		4.05-100.00	672	21.05	0.008	87
ttau	14.29		0.00-85.71	566	23.69	0.004	82
ptau	28.57		0.00-71.43	94	2.63	1	82
Aβ42/ttau	71.43		8.71-100.00	2.093	7.90	0.25	92

Analyte concentrations are described as pg/ml or calculated as ratios between concentrations of two analytes

<sup>a</sup>Cutoff value corresponding to the highest Youden index

<sup>b</sup>Percentage of positively classified cases based on the CSF cutoff value compared with amyloid PET classification

<sup>c</sup>Percentage of subjects who were classified differently based on the cutoff values from fixed specificities compared with the cutoff values corresponding to the highest Youden index

<sup>d</sup>Significance for the "Difference"

Aβ amyloid beta, AUC area under the receiver operating characteristic curve, Cl confidence interval, CSF cerebrospinal fluid, PET positron emission tomography, ptau <sup>181</sup>phospho-tau, SE standard error, ttau total tau





the number of subjects needed to scan to reach a prespecified number of positive cases.

#### Added value of Aß isoforms Aß38 and Aß40

The shorter isoforms A $\beta$ 38 and A $\beta$ 40 on their own had no diagnostic value to discriminate preclinical AD, in line with previous studies in cognitively intact healthy controls [14], and also in clinical AD patients [26]. In the context of preclinical AD, the added value of the A $\beta$ isoforms mainly occurred when used for calculating ratios. The ratio over A $\beta$ 40 performed better than A $\beta$ 42 alone if a high specificity was required (Table 3).

The impact of using A $\beta$  isoforms on the clinical accuracy is linked in part to the context of use. In some studies comparing clinical AD with healthy controls, the ratio of A $\beta$ 42 over A $\beta$ 38 or A $\beta$ 40 improved overall diagnostic accuracy [27, 28], but in others it did not [26, 29]. For the discrimination between clinically probable AD and non-AD dementias, the discriminative value of A $\beta$ 42/A $\beta$ 40 was similar to that of the ratio over ttau and better than A $\beta$ 42 alone [30, 31]. In the MCI stage of the disease, the predictive value for progression to dementia over a 4-year interval was higher for A $\beta$ 42/A $\beta$ 40 (AUC = 0.866) than for A $\beta$ 42 alone (AUC = 0.768) [13]. In our study, A $\beta$ 42/A $\beta$ 40 still allowed acceptable sensitivity for a specificity of 90 %, and more so than A $\beta$ 42 in isolation.

The reason why ratios perform better than  $A\beta 42$  in isolation may be methodological: the normalization procedure may remove a portion of the preanalytical and analytical variability in the measurement of the protein levels that is in itself unrelated to AD. In that case, as better standards become available for Aβ42 measurement, the benefit of using ratios will diminish. Alternatively, the ratio may perform better than  $A\beta 42$  for biological reasons. Many autosomal dominant forms of AD are associated with an increase in the ratio of  $A\beta 42$ over A $\beta$ 40 [32, 33]. Others, such as the Dutch and the Arctic APP mutation, are associated with the inverse effect [32]. If the driving force in the initial phases of sporadic AD is related to disequilibrium between different isoforms rather than the absolute amount of  $A\beta 42$ on its own, this could theoretically explain why the ratio would be better.

#### Ratio of Aβ42 over ttau

For a fixed specificity of 95 %, the highest sensitivity (71 %) was obtained for A $\beta$ 42 over ttau. Generally, ttau is thought to reflect neuronal loss. Adding the separate measurement of a biomarker that increases with the intensity of the neurodegenerative process may enhance specificity because AD is a multidimensional disease [34, 35] so that adding a second dimension (neuronal loss) improves accuracy of classification. The added value of combining A $\beta$ 42 with ttau for the definition of preclinical

AD is in line with the International Working Group IWG-2 criteria for preclinical AD which advocate for the combined use of both  $A\beta42$  and ttau or ptau [3].

#### CSF cutoff value for positive classification

The optimal A $\beta$ 42 cutoff value for the INNOTEST assay was higher than what is commonly applied in clinical practice. Previous studies have also suggested that cutoff values derived from studies in patients with more or less advanced stages of AD versus controls may not be entirely appropriate for distinguishing amyloid-positive from amyloid-negative healthy cognitively intact older adults [14, 36]. This has implications for clinical trials aiming to sensitively select cognitively intact subjects with increased A $\beta$  aggregation [36].

#### Potential study limitations

Our study has some limitations. The sample size was relatively low and the number of amyloid-positive cases was relatively small. Larger studies of preclinical AD will be needed to confirm the estimates of sensitivity and specificity. The low sample size is related to the strict inclusion and exclusion criteria. All subjects were recruited from the community and volunteered for the lumbar puncture purely for research purposes and were informed beforehand that they would not receive any feedback about their proper CSF results. We also applied strict criteria regarding the normality of the neuropsychological test scores. Given the small sample size we were careful to base our conclusions on the most robust findings: we applied strict correction for multiple comparisons and ascertained that our findings were replicable across different assay types and did not depend on small variations of the PET cutoff value within the range of the known test-retest variability of <sup>18</sup>F-flutemetamol PET. For all these reasons we consider our results reliable despite the relatively small sample size, in particular the comparisons between AUC analyses. The repercussions of fixing specificity at 90-95 % on sensitivity have to be interpreted more cautiously: given the relatively low number of true positives, a change in classification of an individual case from positive to negative may lead to a disproportionately large decrease in sensitivity.

A community-recruited cohort is not equivalent to a population-based cohort and could be prone to a selection bias, targeting subjects concerned about their cognition, subjects who were more educated or more mobile, etc. We were careful not to mention memory, cognition, or related terms in our advertisement. The research question at hand, namely the comparison between CSF and PET for the research definition of preclinical AD, is most pertinent for a community-recruited setting: clinical trials targeting preclinical AD will generally not be based on population-based nor on memory clinic-based cohorts but on community-recruited cohorts. There was no evidence for a positive selection bias compared with other community-recruited cohorts. If anything, also taking into account the prior stratification for APOE  $\epsilon$ 4 in our study, our percentage of amyloid-positive cases was lower than in most other community-recruited studies [37]. In a population-based cross-sectional study of cognitively intact adults 50–89 years old, the frequency of amyloid-positive individuals was similar to that in our study [38]. The proportion of subjects who confirmed subjective memory complaints was also not particularly elevated compared with community-based [39, 40] or population-based studies [41].

Our standard of truth was <sup>18</sup>F-flutemetamol positivity based on an autopsy-validated cutoff value. We have previously demonstrated a high concordance between <sup>18</sup>F-flutemetamol and <sup>11</sup>C-Pittsburgh Compound B for the definition of preclinical AD [42]. The autopsy study covered the different Thal stages 1-5 [43]. However, it remains possible, theoretically, that if measured in a population restricted to cognitively intact older adults, the cutoff value for distinguishing moderate to high neuritic amyloid density from sparse to low density may be lower than what is found in a mixed group including patients with advanced dementia along with dementiafree individuals [43]. According to the current study logic, a case who has low A $\beta$ 42 values but a normal <sup>18</sup>Fflutemetamol value would be considered a false-positive. We cannot, however, exclude that this case is in a preclinical state preceding amyloid deposition detectable by PET [14]. In the selection of subjects who have increased risk of amyloid deposition but who have not yet reached the amyloid positivity threshold, there could still be a role for AB isoforms beyond AB42, although this remains to be demonstrated. The specificity required to define preclinical AD based on biomarkers will depend on the type of clinical trial. Different therapeutic strategies may target different preclinical stages of the disease. Our findings are mainly relevant for those trials that target a phase where amyloid aggregation has already occurred and where a marker must be selected, CSF versus amyloid PET.

#### Conclusion

For selection of subjects with increased PET amyloid load, if a high specificity is required, our data support the use of A $\beta$ 42 over ttau rather than using A $\beta$ 42 alone or the ratios to other A $\beta$  isoforms.

#### Abbreviations

Aβ: Amyloid beta; AD: Alzheimer's disease; AIC: Akaike information criterion; APOE: Apolipoprotein E; AUC: Area under the receiver operating characteristic curve; BDNF: Brain-derived neurotrophic factor; CDR: Clinical Dementia Rating; comp: Composite cortical volume of interest; CSF: Cerebrospinal fluid; ELISA: Enzyme-linked immunosorbent assays; MCI: Mild cognitive impairment; MNI: Montreal Neurological Institute; MMSE: Mini Mental State Examination; MRI: Magnetic resonance imaging; NIA–AA: National Institute on Ageing–Alzheimer's Association; PET: Positron emission tomography; ptau: <sup>181</sup>Phospho-tau; ROC : Receiver operating characteristic curve; SD: Standard deviation; SUVR: Standardized uptake value ratios; ttau: Total tau.

#### **Competing interests**

RV has received research grants from Research Foundation Flanders FWO and KU Leuven, has had a clinical trial agreement for phase 1 and 2 study between University Hospitals Leuven and GEHC, has received nonfinancial support from GEHC (provision of <sup>18</sup>F-flutemetamol for conduct of investigator-driven trial free of cost), and has a clinical trial agreement (local principal investigator) between University Hospitals Leuven and Merck, Forum, and Roche. HMJV is an employee of ADx NeuroSciences. JL was an employee of GE Healthcare. The remaining authors declare that they have no competing interests.

#### Authors' contributions

KA contributed to the study concept and design, acquired the data, performed genotyping, performed statistical analyses, interpreted the data, and drafted the manuscript. JS acquired the data, performed genotyping, and revised the manuscript. HMJV analyzed CSF samples and revised the manuscript. JL analyzed neuroimaging data and helped to revise the manuscript. NN interpreted the data and revised the manuscript. KVL contributed to the study concept and design, and revised the manuscript. KP ontributed to the study concept and design, and revised the manuscript. KP analyzed CSF samples and revised the manuscript. KP analyzed CSF samples and revised the manuscript. All authors read and approved the final manuscript.

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